

# Synthetic Control Under Interference: Detecting and Correcting Bias

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Past presentations:

MPSA 2025, Pitt PolSci 2025, MapleMeth2025, PolMeth 2025, NYU 2025, MIT 2025, LAPolMeth 2025

SCM emerged as an important tool for analyzing rare political events:

- **Civil wars:** Coercion, governance, and political behavior in civil war. *Journal of Peace Research*, 2024
- **Polarization:** Partisan Enclaves and Information Bazaars: Mapping Selective Exposure to News. *Journal of Politics*, 2022
- **Far Right:** Do Voters Polarize When Radical Parties Enter Parliament? *American Journal of Political Science*, 2019
- **Religion & Politics:** Government Religious Discrimination, Support of Religion, and Societal Violence in Western Democracies. *Comparative Political Studies*, 2024
- **Political Economy:** From Rents to Welfare: Why Are Some Oil-Rich States Generous to Their People? *American Political Science Review*, 2024
- **Regimes:** The Rush to Personalize: Power Concentration after Failed Coups in Dictatorships. *British Journal of Political Science*, 2023
- **Institutional change:** Comparative politics and the synthetic control method. *American Journal of Political Science*, 2015

When policies, conflicts, or shocks *spill over* to neighboring regions, do we still have valid donor pools under Synthetic Control?

1. Quick SCM & SUTVA Refresher
2. **Detecting interference**
3. **Bias-Correction Toolkit**
4. Simulation Performance
5. Interference in Applied Research
6. German Reunification Re-analysis



# What is the Synthetic Control Method (SCM)?

- Enables inference with a small number (or single) treated units;
- Build a synthetic version of the treated unit as a counterfactual weighting unaffected units.
- Potential outcomes for treated unit:
  - $Y_{1t}^N$ : Outcome in absence of intervention (counterfactual).
  - $Y_{1t}^I$ : Outcome under intervention.
- Treatment effect:

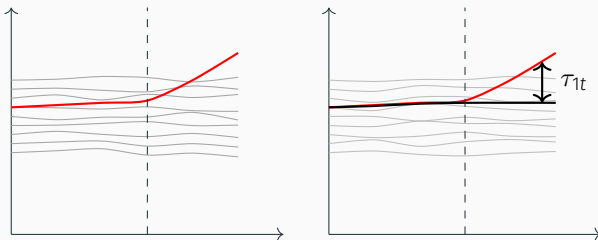
$$\tau_{1t} = Y_{1t}^I - Y_{1t}^N, \quad t > T_0.$$

# SCM: How It Works

$$\hat{Y}_{1t}^N = \sum_{j=2}^{J+1} w_j Y_{jt}, \quad t > T_0.$$

- Optimal weights  $W^*$ : Minimize discrepancy in pre-treatment characteristics and  $\|\cdot\|_V$  reflects predictors importance:

$$W^* = \arg \min_W \|X_1 - X_0 W\|_V,$$



- Stable Unit Treatment Value Assumption (SUTVA):

$$Y_{it}(Z_i, Z_{-i}) = Y_{it}(Z_i) \quad \forall i$$

*No interference: No unit's outcome depends on other units' treatment status.*

- **Crucial Assumption:** The donor units remain *untreated*. Any violation (e.g., partial exposure) can bias the synthetic estimate.
- **SUTVA violation:** Suppose donor  $j$  receives an interference term  $\delta_{jt}$ . The synthetic counterfactual becomes

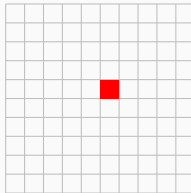
$$\hat{Y}_{it}^N = \sum_{j \neq i} w_j (Y_{jt}^N + \delta_{jt}),$$

so the estimated effect  $\hat{\tau}_{it}$  deviates by  $\sum_j w_j \delta_{jt}$  from the *true*  $\tau_{it}$ .

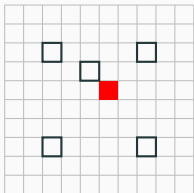
# Stages of SCM Construction



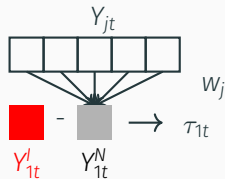
1: Units



2: Single Treated Unit

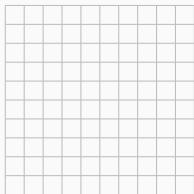


3: Units for Synthetic Control

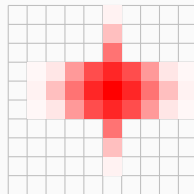


4: Treatment Effect

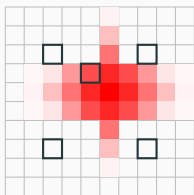
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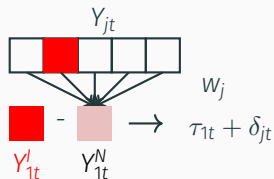
1: Units



2: Treatment diffusion

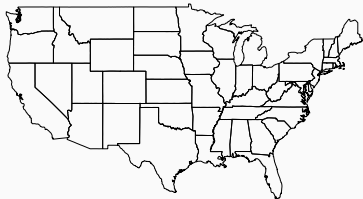


3: Units for Synthetic Control



4: Contaminated Treatment Effect

# Simulated data

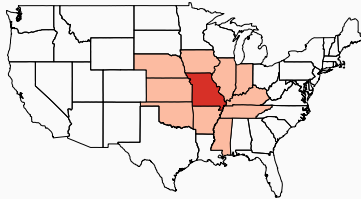


Units map

# Simulated data



Units map

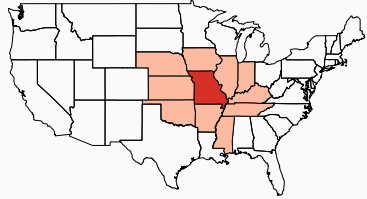


Missouri being treated

# Simulated data



Units map



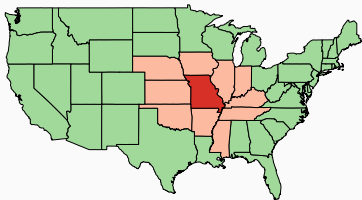
Missouri being treated

Simulated data for an intervention in Missouri with true ATT  $\tau = 4$  and interfering the outcome for nearby units by a parameter of  $\rho = 0.6$

Closer units are more affected by interference than farther away ones. But how can we compare and test if this interference is at play?



# Contrast setup



Contrast for Missouri

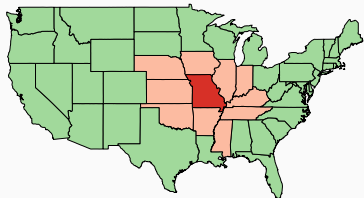
Let  $i \in \mathcal{U} = \{1, \dots, N\}$  index units (in this case, US states)

Fix the treated unit ( $p \in \mathcal{U}$ ) at the center and compute distances  $d_{ip}$  partitioning the space in non-overlapping rings

$$C_0 < C_1 < \dots < C_K$$



# Contrast setup - Z value



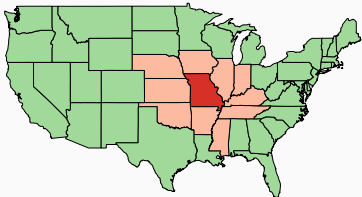
But what are we comparing?

Let  $t \in \mathcal{T}$  index time,  $T_0$  be the treatment period for unit  $p$ , and  $Y_{it}$  represent the outcome

Define two disjoint sets of periods for each window  $w$ :

$$\mathcal{T}_w^{\text{pre}}, \mathcal{T}_w^{\text{post}} \subset \mathcal{T}, \quad \mathcal{T}_w^{\text{pre}} \cap \mathcal{T}_w^{\text{post}} = \emptyset$$

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And set windows of interest for the difference in outcome, such as:

$w$	$\mathcal{T}_w^{\text{pre}}$	$\mathcal{T}_w^{\text{post}}$
full	$\{t < T_0\}$	$\{t > T_0\}$
year-1	$\{T_0 - 1\}$	$\{T_0 + 1\}$
sym- $n$	$\{T_0 - n, \dots, T_0 - 1\}$	$\{T_0 + 1, \dots, T_0 + n\}$

And for every unit  $i$  and window  $w$ , define a difference-in-means statistic:

$$Z_i^{(w)} = \bar{Y}_{i, \text{post}(w)} - \bar{Y}_{i, \text{pre}(w)}$$

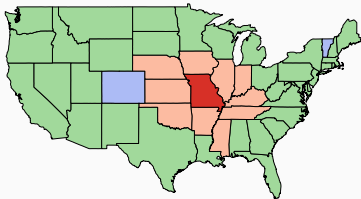
where:

$$\bar{Y}_{i, \text{post}(w)} = \frac{1}{|\mathcal{T}_w^{\text{post}}|} \sum_{t \in \mathcal{T}_w^{\text{post}}} Y_{it}$$

and

$$\bar{Y}_{i, \text{pre}(w)} = \frac{1}{|\mathcal{T}_w^{\text{pre}}|} \sum_{t \in \mathcal{T}_w^{\text{pre}}} Y_{it}$$

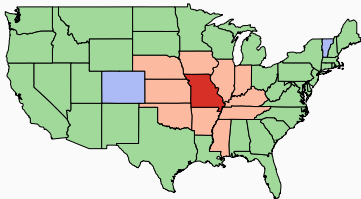
## Contrast setup - first test



$Z_i^{(w)} \rightarrow$  average outcome variation  
for each  $i$  between post-pre periods  
in window  $w$ .

Anomalous values in units nearby  
the treated hint at potential  
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state	$Z^{(\text{full})}$	$Z^{(\text{year-1})}$	$Z^{(\text{sym-3})}$
Missouri	4.0066	3.9159	3.9381
Iowa	2.3640	2.4193	2.3539
Colorado	-0.0414	-0.1069	0.0060
Vermont	0.02501	-0.1115	-0.0886

For each window  $w$ , collect  $Z_i^{(w)}$  for  
 $i \in A_p$  and  $Z_i^{(w)}$  for  $i \in B_p$ , and let


$$\bar{Z}_{A_p}^{(w)} = \frac{1}{|A_p|} \sum_{i \in A_p} Z_i^{(w)}, \quad \bar{Z}_{B_p}^{(w)} = \frac{1}{|B_p|} \sum_{i \in B_p} Z_i^{(w)}$$

denote the group means for each  
ring set and build:

$$t_p = \frac{\bar{Z}_{A_p} - \bar{Z}_{B_p}}{\sqrt{s_p^2 \left( \frac{1}{|A_p|} + \frac{1}{|B_p|} \right)}}$$

Large  $|t_p| \Rightarrow$  evidence that proximity  
ring(s) differ in mean outcome  
change relative to farther rings

## Contrast setup - randomization

Checking whether average  $\neq$  units farther away from   
for nearby units treated unit (around treatment)

Can we reject the null of no interference?

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Can we reject the null of no interference?

**Randomization inference:**

$H_0 : \left\{ Z_i^{(w)} \right\}_{i \in \mathcal{U}}$  is invariant to which unit is labelled “treated”.

i.e.: Pattern of interference around treated unit is no different than the pattern around any other unit in the space



# Contrast setup - randomization II

## Algorithm

1. Compute  $t_p$  for every  $p \in \mathcal{U}$  as above.
2. Let  $t_0$  be the statistic for the **actual treated unit**  $p = p^*$ .
3. Exact two-sided  $p$ -value:

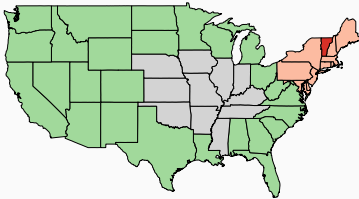
$$\hat{p} = \frac{1 + \sum_{p \in \mathcal{U}} \mathbf{1}(|t_p| \geq |t_0|)}{N + 1}$$

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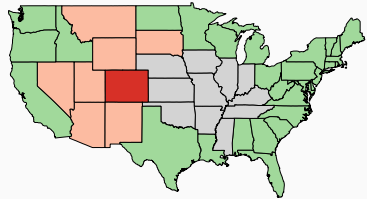
Contrast for Vermont

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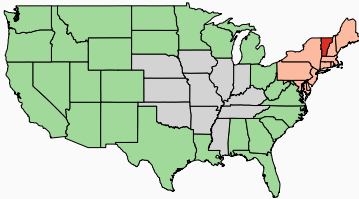
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Contrast for Colorado



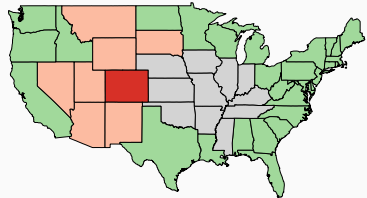
Contrast for Vermont

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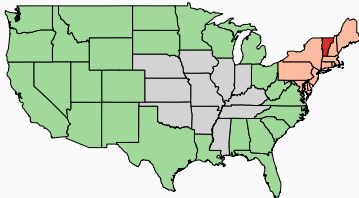
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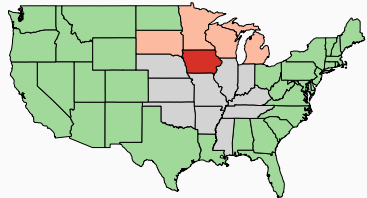
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Contrast for Colorado



Contrast for Vermont



Contrast for Iowa

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state	$t_p$	$A_p$	$B_p$
MO	4.4207	AR, IL, IN, ...	AL, AZ, CA, ...
VT	-0.2169	CT, DE, ME, ...	AL, AZ, CO, ...
CO	0.3428	AZ, MT, NV, ...	AL, CA, CT, ...
IA	-0.3312	MI, MN, SD, ...	AL, AZ, CA, ...

And from this simulated scenario  
we obtained  $p$ -value = 0.0408

## Contrast setup - alternative contrasts

**Where does it end?**

Detecting whether interference is present ✓

Detecting where interference is no longer statistically significant:

# Contrast setup - alternative contrasts

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Detecting where interference is no longer statistically significant:

Instead of contrasting

$A_{p^*} = \{i \neq p^* : r_{ip^*} = 1\}$  vs.

$B_{p^*} = \{i \neq p^* : r_{ip^*} \in \{2, 3, 4, 5\}$

to obtain the standard  $t_{p^*}^{(1 \text{ vs } 2:5)}$

Contrast:  $A_{p^*} = \{i \neq p^* : r_{ip^*} = 2\}$  vs.

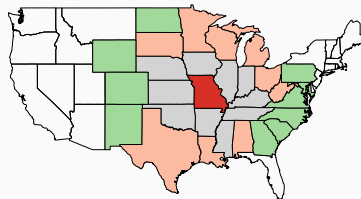
$B_{p^*} = \{i \neq p^* : r_{ip^*} \in 3\} \rightarrow t_{p^*}^{(2 \text{ vs } 3)}$

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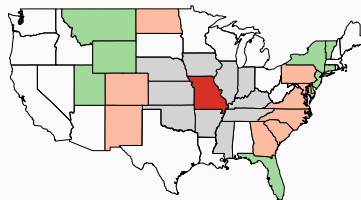
## Where does it end?

Detecting whether interference is present ✓

Detecting where interference is no longer statistically significant:



2 vs 3 Contrast for Missouri,  $p = 0.9591$



3 vs 4 Contrast for Colorado,  $p = 0.5102041$

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## Interference ✓

Two options:

- 1. *Keeping them unmodified* leads to biased synthetic estimates.
- 2. *Simply dropping* suspect donors might degrade the pre-treatment match.

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- 3. *Adjust for it*: Use a secondary set of weights to attenuate contamination in the donor pool

Spatial reach measure as the weights

# Spatial Reach: A Continuous Proximity Index

- For donor  $j$ , let  $d_j$  be its distance to the treated unit.

$$SR_j = \frac{1}{1 + \exp[-\kappa(d_j - c)]},$$

- $c$  is typically the *mean* or *median* distance to center the logistic curve.
- $\kappa$  scales how steeply  $SR_j$  transitions from near 0 to near 1.
- **Parameter Tuning:**  $\kappa$  trimmed between the 2.5% and 97.5% percentiles of  $\{d_j\}$ , ensuring a smooth but complete range.
- **Interpretation:**  $SR_j \approx 0$  if donor  $j$  is very close, and  $\approx 1$  if it is far.

# Bias Correction Strategies

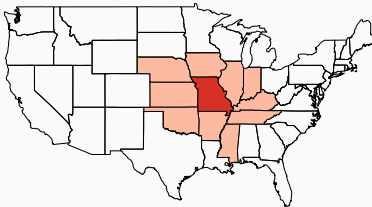
Solution	Optimization	Simplex	Consequence
Rescaling	$\min_w \ X_1 - X_0^* w_j\ ^2$ with $X_{k,j}^* = X_{k,j} \times SR_j$	✓	Downweights exposed units; Retains convex weights
Ridge constrained	$\min_w \ X_1 - X_0 w_j\ ^2 + \lambda \sum_j SR_j w_j^2$	✓	Penalize large SCM weights Moderate contamination
Ridge unconstrained	$\min_w \ X_1 - X_0 w_j\ ^2 + \lambda \sum_j SR_j w_j^2$	✗	Allows negative SCM weights Aggressively offset contamination

Simplex constraint:  $w_j \geq 0$ ,  $\sum_j w_j = 1$

- Units are only allowed to have positive weights
- Unit weights add up to 1

# US Simulation

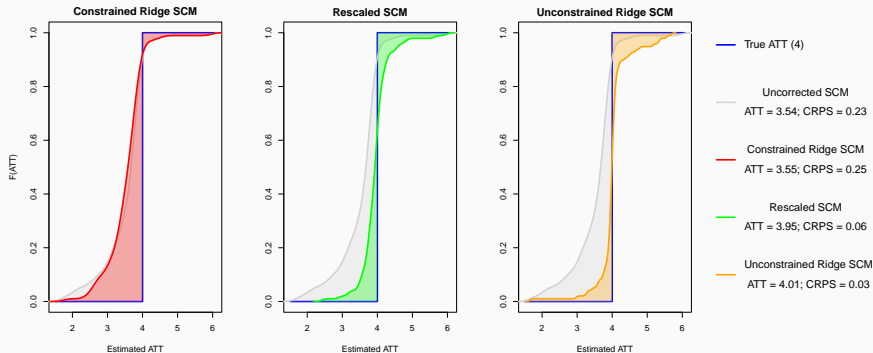
**Setup:** Intervention in Missouri with true effect size  $\tau = 4$  and spillover intensity  $\rho = 0.6$ .



Compare the uncorrected biased SCM versus the three correction approaches

**Metrics:** Bias in the estimated ATT, pre-treatment RMSE, and CRPS.

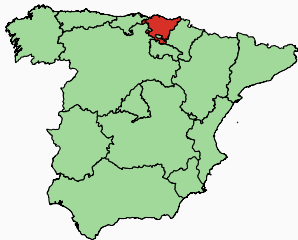
# US Simulation results



Simulation under  $\tau = 4$  and  $\rho = 0.6$

Consistent across all effect sizes  $\tau$  and spillover intensity  $\rho$

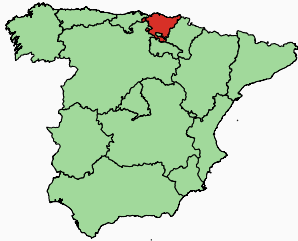
# Interference in Applied Research



Abadie et al (2003) Conflict in the Basque  
 $p = 0.22$



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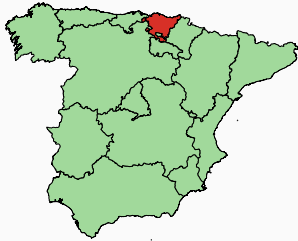


Abadie et al (2003) Conflict in the Basque  
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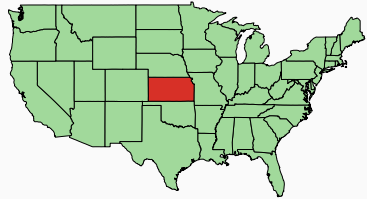


Abadie et al (2015) German Reunification  
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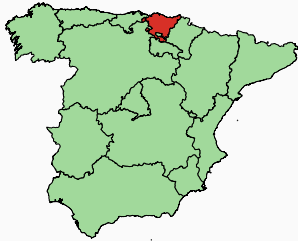


Ben-Michael et al (2021) Kansas tax cut  
 $p = 0.18$

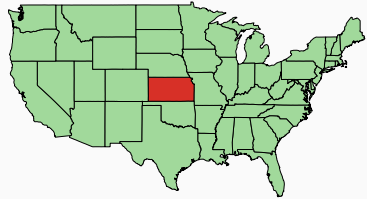


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Abadie et al (2015) German Reunification  
 $p = 0.46$



Kikuta (2020); Civil war and deforestation  
 $p = 0.33$

# Interference in Applied Research

Application	Coverage	Interference
Abadie et al (2003)	✓	✗
Ben-Michael et al (2021)	✓	✗
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Expanded German Reunification	✓	✓



Abadie et al (2015) German Reunification  
 $p = 0.46$



Expanded German Reunification  
 $p = 0.016$

# Interference in Applied Research

Researchers try to address SUTVA violations and patterns of interference by removing units → results conditioned on contagion

Risk → dropping too many units

Under Potential Outcomes, the DGP and a suitable identification strategy depends on: empirics *AND* how the missing potential outcome is set up

- In the SCM case: which units are in the donor pool



# Replication Examples

**Comparative politics and the synthetic control method** (Abadie, Diamond, & Hainmueller, 2015): German Reunification

Approach	Metric	Germany
Base	ATT	-1549.9
	Pre-RMSE	119.08
Rescaled	ATT	-1601.5
	Pre-RMSE	279.03
Penalized, Constrained *	ATT	-1103.4
	Pre-RMSE	80.43
Penalized, Unconstrained *	ATT	136.1
	Pre-RMSE	59.5

*Rescaling* adjusted for contamination → larger effect

*Constrained Ridge* adjust for contamination and large weights → attenuation

*Unconstrained Ridge* extrapolate simplex for aggressive correction → reversal

### A) Detection

- **Coverage:** Ensure proper donor units coverage to compose the missing potential outcome;
- **Detection test:** Using randomization inference, assess whether interference is at place in the empirical setting;
- **Alternative contrast:** By adapting the contrast, identify where interference is no longer detected;
- **Detect Interference First:** If no violation is detected, standard SCM suffices;

### B) Correction

- **SR weight:** If interference  $\rightarrow$  subject the SCM optimization problem to network-specific weights;
- **Minor to moderate interference:** Rescaling or Constrained Ridge can mitigate moderate bias while retaining the notion of a convex combination.;
- **Severe Interference:** Unconstrained Ridge achieves lower bias at the cost of extrapolating out of the simplex;

# Ongoing Extensions

- Inverse Propensity Weighting for Rescaling Approach

HT-Hájek Spatial Weights

Spatial-reach  $f(d)$  as *propensity to avoid spillover*:  $\pi_i = 1 - f(d_{iD})$

Use stabilized Horvitz-Thompson weights

$$w_i = \frac{1/\pi_i}{\sum_j 1/\pi_j} \text{ inside SCM}$$

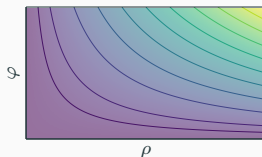
- Multiple Comparison & Dynamic Networks

- Sensitivity to Interference

Inject controlled spillovers in outcomes & covariates: intensity  $\rho \in [0, 1]$ , decay  $\varphi$

Re-run SCM over a  $(\rho, \varphi)$  grid; track standardized shift

Contours show ATT shift required to overturn conclusions



(Lighter  $\rightarrow$  larger ATT shift)